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A new measure of feeling safe: Developing psychometric properties of the neuroception of psychological safety scale (NPSS)

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Conflict of interest statement

There are no known conflicts of interest for any of the authors

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**Data availability statement**

Anonymised data available on request due to privacy/ethical restrictions

**Key words:** Psychological safety, Polyvagal Theory, neuroception, psychometric validation, measure development, trauma, post-traumatic growth, mental health.
A new measure of feeling safe: Developing psychometric properties of the neuroception of psychological safety scale (NPSS)

Objective: Psychological safety is increasingly recognised as central to mental health, wellbeing and post-traumatic growth. To date, there is no psychometrically supported measure of psychological safety combining psychological, physiological and social components. The current research aimed to develop and establish the neuroception of psychological safety scale (NPSS), informed by Polyvagal Theory.

Method: The study comprised of three stages: (1) item generation, (2) item reduction, and (3) assessment of factor structure and internal consistency. Exploratory and confirmatory factor analysis was conducted from two samples who completed a survey online (exploratory n = 342, confirmatory n = 455).

Results: Initially, 107 items were generated. Item reduction and exploratory factor analysis resulted in a 29-item NPSS with subscales of compassion, social engagement and body sensations. The NPSS was found to have a consistent factor structure and internal consistency.

Conclusion: The NPSS is a novel measure of psychological safety which can be used across a range of health and social care settings. This research provides a platform for further work to support and enhance understandings of the science of safety through the measurement of psychological, relational and physiological components of safety. The NPSS will help shape new approaches to evaluating trauma treatments, relational issues and mental health concerns. Research to establish the convergent, discriminant and concurrent validity of the NPSS and to explore its use with diverse community and clinical populations is underway.
Clinical Impact:

Psychological safety is recognised as central to mental health, wellbeing, post-traumatic growth, therapeutic engagement and an understanding of trauma related conditions and trauma informed practices.

The Polyvagal Theory (PVT) offers a comprehensive explanation of psychological safety grounded in an evidence-base of neurophysiology, psychology and evolutionary theory and the current study aimed to develop a self-report measure, the Neuroception of Psychological Safety Scale (NPSS) informed by the PVT.

The NPSS integrates psychological, relational and physiological components of safety and has applicability across a range of health and social care contexts, such as shaping new approaches to evaluating trauma treatments as well as broader relational issues and mental health concerns.

Introduction

Psychological safety is recognised as central to mental health, wellbeing (Sullivan et al., 2018) and post-traumatic growth (Norman et al., 2020) with increasing clinical interest and research attention toward its importance. Feeling safe is recognised as a distinct state important for rest, restoration and social bonding (Goetz et al. 2010; Porges, 1998). As social beings perceived threat is often interpersonal while safety with other people is communicated using compassion (Gilbert, 2017). Compassionate interventions, such as the use of soothing voice tones and breathing, reduce the fight/flight response, decelerate heartbeat and facilitate parasympathetic rest and restoration (Kirby et al., 2017). A safe and compassionate early environment shapes the nervous system and aids the development of self-soothing strategies that enable self-regulation in
later life (Gilbert, 2017). Trauma symptoms arise from unregulated threat preoccupation, when self-regulation is not available, which affects our biology, social interaction, and maturation (Motsan et al, 2021; van der Kolk, 1994).

To date, psychological safety research has largely been considered within organisational and group contexts, describing the process of assessing risk in interpersonal relationships and occupational environments. The Team Psychological Safety Scale (Edmondson, 1999) is a 7-item self-report scale that measures perceptions of feeling safe within teams which has good reliability and validity (Chen et al., 2015). Increased sense of psychological safety at work facilitates employee communication, improvements in learning, teamwork and work performance (Edmondson and Lei, 2014; O’Donovan et al, 2020). The positive impact of psychological safety has been found in other organisational contexts, including public spaces, education (Wanless, 2016), community building (Singh et al., 2018), virtual meetings (Edmudson and Daley, 2001), communicating in medical teams (Real et al, 2021) and in healthcare workplaces to reduce levels of psychological distress and trauma (Ahmed et al., 2021). However, psychological safety and its measurement differs within teams differs from the individual.

Psychological safety for the individual, rather than within teams, has also begun to gain attention within mental health settings regarding clinical understanding of trauma related conditions and trauma informed practices (Isobel et al., 2021) where traditional measures focus on pathology rather than prevention and positive adaptation. Difficulty in assessing danger or safety and modulating fear response is reported in individuals suffering Post Traumatic Stress Disorder (PTSD) (Jovanovic et al., 2012). A novel manualised cognitive-behavioural treatment for PTSD called ‘Seeking Safety’ which prioritises feeling safe (Najavitis, 2001) delivered
improved outcomes in symptoms of PTSD and psychiatric distress compared to controls (Desai et al., 2008). ‘The Feeling Safe Program’ aims to address safety feelings when treating persecutory delusions in psychosis and a clinical trial of this intervention showed recovery (Freeman et al., 2016).

The following psychological scales include a component of psychological safety. In the Activation and Safe/Content Affect Scale (Gilbert et al., 2008) safe affect is shown to negatively correlate with measures of depression, anxiety, stress, self-criticism, and insecure attachment. The same research team developed the Scale of Childhood Memories of Emotional Warmth and Safety (Richter et al., 2009). The Therapeutic Environment Scale includes a ‘feeling safe with others’ subscale, validated using clinical samples (Veale et al., 2016). The Child Safety Behaviour Scale has been developed to measure safety-seeking behaviours (Alberici et al., 2018) but is less concerned with affective states.

In medical settings concern for the sense of safety experienced by patients (Ellegaard et al., 2020; Morton 2020) and when exposed to disempowering aspects of care (Morton et al., 2019) is of interest in terms of quality of experience and speed of recovery. In one study, feeling safe during the process of hospitalisation was found to increase feelings of control, calm and hope (Mollon, 2014). Feeling safe has also been found to improve healing and recovery during maternity care of women who have experienced childhood sexual trauma, whilst feeling unsafe with professionals could be experienced as re-traumatisation (Montgomery, 2013).

In sum, to date psychometric measures of feeling safe have been restricted to specialised contexts such as team safety (Edmondson, 1999) childhood memory of safety (Richter et al., 2009), as a subscale (Veale et al., 2016), or as a dimension of a broader scale under factor analysis (Gilbert et al., 2008) rather than the central construct. Due to the importance of safety
within a therapeutic context and the lack of a general dedicated means of measurement that considers relationship dynamics (Roussin et al., 2016), there is a need for the development of a refined psychometrically validated scale of psychological safety.

The Polyvagal Theory (PVT) offers a comprehensive explanation of psychological safety grounded in an evidence-base of neurophysiology, psychology and evolutionary theory. PVT describes how situations are subconsciously assessed for safety or threat by the autonomic nervous system, termed ‘neuroception’, leading to corresponding physiological, affective, and behavioural responses (Porges, 2004). In developing a scale of psychological safety, PVT proposes that situations detected via ‘neuroception’ as safe will activate physiological, affective, and cognitive processes to optimise social engagement through compassion for others. Situations detected as unsafe will shift bio-behavioural systems that would restrict interpersonal social engagement, while optimising physiological state, via the autonomic nervous system to support defensive survival strategies either via the dorsal vagal pathway leading to immobilising, death feigning, or dissociating or via the sympathetic system leading to fight/flight behaviours that would be supported by increases in heart rate, shortened breathing, and increased muscle tension (Kolacz et al., 2019).

PVT has helped to inform mental health, medical, and educational practices in the use of safe therapeutic presence (Geller and Porges, 2014), recognition of client’s non-verbal safety-signalling (Mair, 2020) and interpreting representations of fear and safety in art therapy (Gerge, 2017). It has also acted as the basis of the Body Perception Questionnaire (Cabrera et al., 2018) and the Brain-Body Center Sensory Scales (Kolacz et al., 2018) supporting the utility of PVT as the grounding for a new general scale of safety. As such, the current study aimed to develop
such a self-report measure, the Neuroception of Psychological Safety Scale (NPSS) informed by the PVT (Porges, 2004, 2011, 2021).

**Methods**

**Ethics**

Ethical approval for this research was sought and received from the University Ethics Committee and all participants provided informed consent to their participation prior to engaging with the research. All participants were recruited using convenience sampling using online platforms, Facebook, Twitter as well as university recruitment posters.

**Phase 1: Development of statement items**

Item development for this study followed the best practices for domain identification, item generation, and theoretical analysis (Boateng et al, 2018). A variation of the e-Delphi methodology was employed to facilitate the first phase of the process Galanis, 2018; Skulmoski, Hartman, and Krahn, 2007).

Six international experts in the PVT, Clinical and Counselling Psychologists, and therapists, with academic and clinical experience in trauma work, were identified. These experts were recruited to include leading expertise in the PVT (SP, founder of the PVT; Deb Dana, Psychotherapist, founding member of the Polyvagal Institute and leader in PVT informed therapies; JK), development of PVT-informed psychometric measurement (JK), and clinical and academic expertise in trauma (NC, Consultant Clinical Psychologist / Senior Lecturer; LM, Counselling Psychologist/ Lecturer; EP, Counselling Psychologist/Lecturer; TB, Consultant Clinical Psychologist).
Applying the e-Delphi method (Galanis, 2018) the stakeholders first agreed on the aims of the study which were to generate a consensus on a list of items pertaining to psychological safety. To generate items each expert stakeholder anonymously provided items pertaining to what it means to feel safe via email. Specifically, in the first round an open-ended questionnaire asking ‘What does it mean to feel safe?’ was used to systematically solicit and collect items from all stakeholders, anonymously and individually, via email to the lead author (LM). Following this item generation, these items were collated by the lead author who re-distributed the complete list via email to each of the key stakeholders inviting feedback and critique and the opportunity to re-evaluate their initial responses. Subsequently, there was a process of combining and redistributing the items until a consensus over a final set of items was reached by all stakeholders. Following four rounds 107 items were agreed by all stakeholders.

**Phase 2: Item reduction and piloting**

The second phase of developing the NPSS involved piloting the items generated in Phase 1 to reduce the number of items and to ensure that the questions were relevant, clear, and measured what they were intended to measure (Boetang et al., 2018).

**Procedure**

Firstly, this was achieved through assessing the 107-item NPSS and removing items that did not contribute to the measure using stakeholder review and exploring the factor loadings of the statement items. The 107 items formed the basis of the piloting of NPSS. The stimulus for the NPSS used to evaluate the items was “Think about a recent specific situation when you felt safe, and please rate the following statement items in relation to this.” The final set of the items was randomized in accordance with a five-point Likert scale, with lower scores indicating less
agreement with the item e.g., *Strongly Disagree* (score = 1), *Disagree* (score = 2), *Neither Agree or Disagree* (score = 3), *Agree* (score = 4), *Strongly Agree* (score = 5). Two other options were added *Prefer not to say* and *Not applicable*. Collected data were treated as ordinal as appropriate of data regarding attitudes (Hair *et al*, 2019). Statistical analyses were conducted with RStudio software for Windows, version 1.2.5019 (RStudio, Inc., 2009-2016), and IBM SPSS (IBM SPSS, Inc., 2017). Factor analysis was completed employing the fa function included in Psych package for R (Revelle, 2017).

**Exploratory Factor Analysis**

An exploratory factor analysis (EFA) of the NPSS was conducted to examine the factor structure and reduce items. The exploratory factor retention was guided by factor loading, theoretical predictions, scree plots, and model fit. The goodness of fit to the data was evaluated by using the root mean squared error of approximation (Steiger, 1990), Tucker-Lewis index (Tucker and Lewis, 1973), and the Comparative Fit Index (Bentler, 1990). Eigenvalues were observed for the last considerable drop in magnitude. The factor analysis was subject to oblimin rotation; factors related to safety were expected to correlate and not to be highly complex (Fabrigar *et al*., 1999). Factoring method by minimal residuals was used (minres), discussed to be optimal for oblique data (Brown, 2009). Internal reliability was assessed using the Cronbach’s $\alpha$ coefficients with the threshold for acceptability at 0.7 (Nunnally and Bernstein, 1994). In support, Omega $h$ was calculated to assess the lowest level of reliability to multiple latent factors. Dunn *et al* (2014) summarise the advantages of omega over alpha. Only items loading > .6 were retained unless cross-loaded with another factor >.45 (Ladhari, 2010).

**Results**

342 participants consented to and completed the NPSS. Data were accepted for inclusion if most of the NPSS was completed and participants did not fail attention checks. Six participants
were excluded based on failing attention check. Attention checks required participants to provide a specific answer (e.g., “Please, select Not-applicable”). The final number of responses for analysis was 336 (229 women, 103 men, 4 non-binary). The mean age of participants was 32.87 (SD 12.17, range 54). Ethnicity was predominantly given as white Scottish.

Missing data were deleted pairwise, so if a participant completed most but not all the NPSS, their answers were included in the analysis. A correlation matrix was calculated. A single measure of correlation strength was obtained by calculating a back-transformed mean Fisherman Z score for each item (Salkind, 2013). 33 items were removed prior to factor analysis for the following reasons: 1. Duplicate items and very similar items were reduced to a single item. Correlations among the pairs and groups of similar items were explored. If their face similarity was supported by correlation >.7, only the highest correlating item remained. 2. Items that correlated poorly with all other items were deleted. The final number of items evaluated in the EFA was 74.

**Factor structure of the NPSS**

Polychoric correlations were calculated and a polychoric parallel analysis was employed to explore the factor structure as appropriate for ordinal data (Garrido *et al.*, 2013).

Decisions for factor retention were guided by a combination of absolute fit indices, scree plot, and simple structure (Montoya & Edwards, 2021, Fabrigar *et al.*, 1999). The NPSS scree plot and goodness of fit indices did not provide a clear solution on the first iteration. Visual examination of the scree plot (see supplementary materials) indicated one large eigenvalue followed by relatively small values without a clear second drop, indicating a 1-factor solution. One factor solution was not supported by model fit statistics (RMSEA = .093 [90%CI: .091-.095], CFI = .63, TLI = .62). On second iteration, items that did not load substantially on any
factor were dropped. The resulting pool was reanalysed. The visual analysis of the scree plot showed three deviating eigenvalues suggesting a 3-factor model to explain the data (Table 1.).

**INSERT TABLE 1 HERE**

Items that did not load substantially on any factor were dropped. The resulting item pool was reanalysed. The visual analysis of the scree plot showed three deviating eigenvalues suggesting that 3-factor model could be used to explain the data (see supplementary materials). All items demonstrated simple structure by loading substantially on only 1 factor, thus we accepted the 3-factor solution ($RMSEA = .083 [90\% CI: .080, .086]$, $CFI = .80$, $TLI = .78$).

Bartlett's Test of Sphericity indicated a good factorability ($\chi^2 = 398.96; df = 73; p < .01$) of the correlation matrix. The correlations between the factors were moderate; between the first- and the second-factor $r = .65$, between the first- and third-factor $r = .52$, and between the second- and third-factor $r = .50$. Factor loading thresholds of loading $>.6$ were applied (Wolfinbarger and Gilly, 2003). The resulting factors corresponded with the theoretical concepts of safety; Compassion, Social Engagement and Bodily Sensations. All items within a factor conceptually coalesce. Items demonstrated a simple structure, and none of the items cross-loaded substantially on more than 1 factor. Resulting items and factor loadings are presented in Table 2. Cronbach’s $\alpha$ and Omega $h$ were calculated for each subscale separately.

**INSERT TABLE 2 HERE**

**Internal consistency**

Descriptive statistics for NPSS are presented in Table 3. The NPSS deviated from normality as assessed by skewness, kurtosis, and Shapiro-Wilk tests ($p < .05$). Reliability measures used do not rely on normality assumptions. Internal consistency was assessed by
Cronbach’s $\alpha$ and omega hierarchical coefficient. Cronbach’s alpha has been shown robust to non-normally distributed data when the dataset is sufficiently large, but it can be affected by test score distributions (Sheng and Sheng, 2012). Omega hierarchical, implemented in the Psych R package (Revelle, 2009) provides superior internal consistency assessment when data is ordinal. Omega $h$ scores range from 0-1, appropriate values for psychometrically evaluated scales ranged from 0.6 to .98 (Viladrich et al., 2017).

**Post Hoc Analysis**

As described the resulting factors corresponded with the theoretical concepts of safety; Compassion, Social Engagement and Bodily Sensations which were considered subscales. A total score of each subscale was calculated for each participant to analyse the difference between genders (Norman, 2010). Independent sample $t$ test revealed that there was no significant difference between genders on Social Engagement and Compassion subscales. Males ($n = 92$, mean = 31.74, SD = 5.54) had significant lower scores than females ($n = 217$, mean = 33.2581, SD = 4.36) on the Body Sensations Subscale after correction for unequal variances ($t (140.9) = 2.340, p = .021$, Cohen’s $d = .3$, equal variances not assumed). Similar results were obtained after 1000 bootstrap samples ($p = .026$) to control for unequal sample sizes.

**Phase 3: Internal Reliability and Dimensionality**

**Procedure**

The third phase involved administering the 29-item version of the NPSS to a sample of participants using the same inclusion criteria and procedure as in phase 2. The aim was to assess the internal reliability and dimensionality of the NPSS. Participants were invited to complete the 29 item NPSS. Five demographic questions were included; gender, age, ethnicity, country of
residence and working status. The present study utilised a survey design for the purpose of psychometric evaluation. In tests of dimensionality and reliability, the NPSS was independently assessed at item-level. Monitoring of data part-way through collection revealed a low male response rate. Targeted recruitment on aforementioned social media sites was conducted using an updated recruitment poster. A total 455 participants completed the survey. Excluding participants who did not fill the NPSS and standardised measures or otherwise showed excessive or non-random missingness on these measures, this number was reduced to 318 (86 males, 5 non-binary or not-sure). The mean age of participants was 22.66 (SD = 14.65, range = 18-71). Ethnicity was predominantly given as white Scottish.

**Analysis**

Statistical analysis was conducted using RStudio software for Windows, Version 1.3.1093 (RStudio inc., 2009-2020). In order to maximise available data and reduce biased estimates, an imputation method was applied to missing item data. As data appeared to be missing at random and rates of missingness were low (1.13%) a single imputation method was deemed appropriate (Dong and Peng, 2013). This was conducted using the expectation-maximisation algorithm. Descriptive statistics for the NPSS were generated, giving indications of normality and potential outliers. A content analysis was conducted for answers to the NPSS open-ended stimulus-prompt question, uncovering themes and frequency of their occurrence in participant-generated situational stimuli. Dimensionality of the NPSS was measured using confirmatory factor analysis (CFA), with predicted factors based on an exploratory factor analysis conducted during the first phase. Cronbach’s Alpha was calculated to determine internal reliability of the scale and sub-scales in the NPSS. Whilst Cronbach’s Alpha is a standard test of reliability, it has been found to show bias for scales that are not unidimensional or that do not show tau-equivalence, and so McDonald’s Omega was also conducted to triangulate results
Sum scores of the NPSS were calculated and used in validity testing. Known groups validity was investigated via an independent t-test comparison of men and women’s NPSS scores, predicting significantly lower results for women in line with published literature (Logan and Walker, 2017).

**Results**

Initial data exploration uncovered a sum score mean of 119.48, (SD 14.85) and item mean of 4.12. Negative skew (-0.97) and positive kurtosis (3.75) were found in NPSS sum scores. A meaningful pattern in low-scoring outliers of the NPSS was also identified. In answer to the situational prompt open-ended question, these individuals chose asocial situations such as being home alone, unlike the majority of participants who chose a social situation. Removal of outliers showed no change in non-normality and so were included in dimensionality and reliability testing. Outliers were deleted from correlation and regression analysis however, as these tests are sensitive to the presence of outliers and may also lead to reduced linearity of variables (Wilcox, 2016).

**Open-Ended Question Content Analysis**

A total 86 responses were given for the open-ended situational prompt question. A conceptual content analysis revealed that 58 (67.44%) chose interpersonal situations involving loved ones, friends, colleagues or caring professionals. Only 3 (3.48%) chose explicitly asocial situations, whilst the remainder focused on location, situation or material safety (such as being at home or at work).

**Dimensionality**

The data were assessed for multivariate normality using Mardia’s test. Due to failure of testing, factor analysis results are reported with robust standard errors and Satorra-Bentler corrections. Fit indices show that the proposed 3-factor model met recommended cut-off levels.
of good fit for X2 (difference X2 = 772.44, df = 374, p<.001), RMSEA (0.058) and SRMR (0.062), whilst CFI (0.86) and TLI (0.84) closely approached cut-off levels (Hu and Bentler, 1999). Factor loadings show acceptable fit of items to the three factors, however question 1 ‘I felt valued’, 2 ‘I felt comfortable expressing myself’ and 14 ‘I didn’t feel judged by others’ did not meet stricter standardised loading cut-offs of ≤ 0.7 (MacCallum et al., 2001).

**Internal Reliability**
Cronbach’s α for the entire NPSS was .95. Subscale results were .93 for social engagement, .94 for compassion and .92 for bodily sensations, suggesting good internal reliability of the scale and subscales (Agbo, 2010). In all cases, α could not be increased by excluding items. Omega h total scores were .96 for overall NPSS, .93 for social engagement, .93 for compassion and .92 for bodily sensations, confirming the findings of Cronbach’s α testing (McNeish, 2018).

**Discussion**

The study reports that the NPSS is a psychometrically sound measure that captures the multiple dimensions of psychological safety that people experience but that, until now, have been difficult to operationalise and measure. The first phase resulted in generation of 107 items pertaining to what it means to feel safe by psychologists and researchers with expertise in trauma and the PVT creating the comprehensive NPSS. The second phase evaluated the items and assessed factor structure, thus creating the 27-item NPSS scale with three subscales consistent with understanding of safety as proposed by the PVT and literature in psychological safety. The first factor, termed Social Engagement, is characterised by being accepted, understood, cared for, being able to express oneself without being judged, and having someone to trust. The items indicated evaluation of the social environment as non-threatening and safe to engage socially – a property ascribed to the Social Engagement System (SES) (Porges, 2011). The second factor
captured items related to the ability to be compassionate and feeling connected, empathetic, caring and wanting to help. Being compassionate regulates our autonomic nervous system (Kirby et al., 2017) while regulation occurs through the ability to self-soothe (Mok et al., 2019) and communicating safety. In therapy, compassion is increasingly seen as central to promote safety and develop/re-engage self-soothing strategies (Gilbert, 2017). The third factor related to the internal sensations of the body in a state of calm capturing the feeling of relaxation in the face and the body, steady heartbeat and breath, and settled stomach. The activation and functioning of the SES are associated with the regulatory function, especially of the heart and bronchi and the associated state of relaxation and restoration (Porges, 2011).

Correlation was stronger between the first and second factors which may suggest a bidirectional link between the feeling of being accepted within a group and compassion (Liu, 2017). We found a gender difference on the Body Sensations subscale but not on the other two subscales with males scoring significantly lower. Body awareness has also shown to be impacted by age (Cabrera et al., 2018) and psychopathology (Bernatova and Svetlak, 2017), which may be considered in further evaluative efforts.

In the third phase the NPSS was evaluated with CFA. The three-factor structure showed adequate fit, and the scale showed good reliability. In both phase two and three, scores distribution was leptokurtic and negatively skewed. This ceiling effect observed is attributed to sampling from the convenience general population sample and participants being prompt to ‘think about a recent specific situation when you felt safe’ which may have led to participants responding about a situation when they felt optimally safe.

Applicability of the NPSS may be improved by inquiring about a specific situation or event, e.g., ‘please rate the following statements in relation to [insert your event]’. However, the
original wording may be useful when determining the base-line safety is desirable, for example when the objective is mental health recovery (Lewis et al., 2019). The applicability of the NPSS could also be increased by formalising a procedure for scoring and interpreting subscale measures, as the bodily sensations subscale may be more useful in gauging feelings of safety in asocial situations. Though the NPSS is a relatively brief instrument, future studies are needed to explore whether a shorter form may be developed to expand clinical utility for cases in which rapid measurement is a priority.

This study has several limitations and suggestions for future work. First, since participants self-selected through convenience sampling rather than being randomly selected, they may have had stronger feelings about psychological safety than those in the larger population in general. While recruiting participants through social media has many advantages, it also has its potential biases, that may limit generalisability (Benedict et al., 2019). Future directions aim to understand more about the latent factors underlying the dimensions of the NPSS via CFA and once understood, will explore the possibility of creating a briefer measure. We are also collaborating with colleagues in clinical practice, who are engaged in trauma practices, to test the feasibility of using the NPSS in intervention planning and goal setting with service users. Further, future work could explore the psychometric properties and feasibility of the NPSS in more diverse populations and across a range of socio-cultural contexts. While the difference in neurophysiological expression of safety has insofar not been addressed by researchers, cultural differences in perception of safety and risk-taking have been identified (Liao, 2015). Exploring psychological safety within dyadic relationships and mapping these using social network methods will allow for an improved understanding of how individual level factors can be considered with relationships and social structures that shape outcomes (Holman
and Borgstrom, 2016; Roussin et al., 2016). Future research should consider the impact of social networks, systemic factors and cultural contexts and their impact on psychological safety within therapeutic contexts (Ash et al, 2021). The NPSS is a new and novel measure of psychological safety which integrates psychological, relational and physiological components of safety and has applicability across a range of health and social care contexts. The NPSS may shape new approaches to evaluating trauma treatments (e.g., compassion focused therapy) as well as broader relational issues and mental health concerns. For example, using the NPSS to better understand shifts in the window of tolerance of autonomic arousal (Porges, 2011) coupled with further work capturing narratives of what it means to be psychologically safe for clients who have experienced trauma. Research to establish the convergent, discriminant and concurrent validity of the NPSS and to explore its use with diverse community and clinical populations is underway.
References


Version = 1.7.8.


**Table 1. Eigenvalues and percentages of variance associated with each component**

<table>
<thead>
<tr>
<th>Component</th>
<th>Eigenvalue</th>
<th>Percentage of explained variance</th>
<th>Accumulated percentage of explained variance</th>
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<tr>
<td>1</td>
<td>23.055</td>
<td>0.242</td>
<td>0.242</td>
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<tr>
<td>2</td>
<td>3.031</td>
<td>0.169</td>
<td>0.412</td>
</tr>
<tr>
<td>3</td>
<td>2.218</td>
<td>0.164</td>
<td>0.575</td>
</tr>
</tbody>
</table>

**Table 2. The items of the Neuroception of Psychological Safety Scale (NPSS: 29 items) with factor loading and reliability score**

<table>
<thead>
<tr>
<th>Subscale</th>
<th>Item</th>
<th>Loading</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Social Engagement</strong></td>
<td>I felt valued</td>
<td>0.751</td>
</tr>
<tr>
<td><strong>Number of items: 14</strong></td>
<td>I felt comfortable expressing myself</td>
<td>0.697</td>
</tr>
<tr>
<td></td>
<td>I felt accepted by others</td>
<td>0.796</td>
</tr>
<tr>
<td></td>
<td>I felt understood</td>
<td>0.825</td>
</tr>
<tr>
<td></td>
<td>I felt like others got me</td>
<td>0.697</td>
</tr>
<tr>
<td></td>
<td>I felt respected</td>
<td>0.721</td>
</tr>
<tr>
<td></td>
<td>There was someone who made me feel safe</td>
<td>0.675</td>
</tr>
<tr>
<td></td>
<td>There was someone that I could trust</td>
<td>0.684</td>
</tr>
<tr>
<td></td>
<td>I felt comforted by others</td>
<td>0.712</td>
</tr>
</tbody>
</table>

*Cronbach’s α 0.93  
*Omega h 0.81
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<th>Median</th>
<th>SD</th>
<th>Skew</th>
<th>Kurtosis</th>
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<td>NPSS</td>
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<td>4</td>
<td>0.85</td>
<td>-0.92</td>
<td>0.64</td>
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<tr>
<td>Social Engagement Subscale</td>
<td>4.19</td>
<td>4</td>
<td>0.86</td>
<td>-1.03</td>
<td>0.83</td>
</tr>
<tr>
<td>Compassion Subscale</td>
<td>4.17</td>
<td>4</td>
<td>0.82</td>
<td>-0.82</td>
<td>0.44</td>
</tr>
<tr>
<td>Body Sensations Subscale</td>
<td>4.04</td>
<td>4</td>
<td>0.83</td>
<td>-0.77</td>
<td>0.49</td>
</tr>
</tbody>
</table>

Table 3. Descriptive statistics for the NPSS and subscales